

## Common Claim Frequency Distributions

- **Binomial Distribution,  $X \sim \text{binomial}(n, p)$ :**

$$p(x|n, p) = \binom{n}{x} p^x (1-p)^{n-x},$$

where  $x = 0, 1, 2, \dots, n$ ,  $n = 1, 2, \dots$ , and  $p \in [0, 1]$ ,

$$E[X] = np, \text{ and } \text{Var}[X] = np(1-p).$$

If  $X \sim \text{binomial}(1, p)$ , then  $X \sim \text{Bernoulli}(p)$ .

If  $X_1 \sim \text{binomial}(n_1, p)$  and  $X_2 \sim \text{binomial}(n_2, p)$  are independent random variables, then  $X_1 + X_2 \sim \text{binomial}(n_1 + n_2, p)$ .

- **Poisson Distribution,  $X \sim \text{Poisson}(\lambda)$ :**

$$p(x|\lambda) = \frac{e^{-\lambda} \lambda^x}{x!},$$

where  $x = 0, 1, 2, \dots$ , and  $\lambda \in (0, +\infty)$ ,

$$E[X] = \lambda, \text{ and } \text{Var}[X] = \lambda.$$

If  $X_1 \sim \text{Poisson}(\lambda_1)$  and  $X_2 \sim \text{Poisson}(\lambda_2)$  are independent random variables, then  $X_1 + X_2 \sim \text{Poisson}(\lambda_1 + \lambda_2)$ .

- **Negative Binomial Distribution,  $X \sim \text{negative binomial}(r, p)$ :**

$$p(x|r, p) = \binom{r+x-1}{r-1} p^r (1-p)^x,$$

where  $x = 0, 1, 2, \dots$ ,  $r = 1, 2, \dots$ , and  $p \in [0, 1]$ .

$$E[X] = \frac{r(1-p)}{p}, \text{ and } \text{Var}[X] = \frac{r(1-p)}{p^2}.$$

If  $r = 1$ , then  $X \sim \text{geometric}(p)$ .

If  $X_1 \sim \text{negative binomial}(r_1, p)$  and  $X_2 \sim \text{negative binomial}(r_2, p)$  are independent random variables, then  $X_1 + X_2 \sim \text{negative binomial}(r_1 + r_2, p)$ .

## The Normal Approximation

- Central Limit Theorem:

If  $X_1, X_2, \dots, X_n \sim \text{i.i.d. } F(x)$  such that  $\mu = E[X]$  and  $\sigma^2 = \text{Var}[X] < +\infty$  both exist, then

$$\Pr \left\{ \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \leq z \right\} = \Pr\{Z \leq z\} = \Phi(z)$$

as  $n \rightarrow +\infty$ .

- The central limit theorem often allows one to assume that the sample mean,  $\bar{X}$ , is approximately  $N\left(\mu, \frac{\sigma^2}{n}\right)$ , or that the sum  $\sum_{i=1}^n X_i$  is approximately  $N(n\mu, n\sigma^2)$ .

The normal approximation becomes more accurate as the sample size,  $n$ , increases. For many applications,  $n$  should be  $> 25$ .

For a fixed sample size, the accuracy of the normal approximation is generally better for random variables  $X_i$  that are less skewed.

- Continuity Correction for Discrete Distributions:

Suppose that one wishes to use the normal approximation to find

$$\Pr\{a \leq Y \leq b\}, \Pr\{a \leq Y < b\}, \Pr\{a < Y \leq b\}, \text{ or } \Pr\{a < Y < b\},$$

where  $Y$  is a discrete random variable defined on the integers, and  $a$  and  $b$  are both integers.

In modeling  $Y$  as a continuous normal random variable, it is more precise to recognize the discrete nature of  $a$  and  $b$  by either adding or subtracting  $\frac{1}{2}$  from these values when normalizing  $Y$ .

- Let  $E[Y] = \mu$  and  $\text{Var}[Y] = \sigma^2$ . Then:

$$\Pr\{a \leq Y \leq b\} = \Pr \left\{ \frac{a - \frac{1}{2} - \mu}{\sigma} \leq Z \leq \frac{b + \frac{1}{2} - \mu}{\sigma} \right\} = \Phi \left( \frac{b + \frac{1}{2} - \mu}{\sigma} \right) - \Phi \left( \frac{a - \frac{1}{2} - \mu}{\sigma} \right), \text{ or}$$

$$\Pr\{a < Y < b\} = \Pr \left\{ \frac{a + \frac{1}{2} - \mu}{\sigma} \leq Z \leq \frac{b - \frac{1}{2} - \mu}{\sigma} \right\} = \Phi \left( \frac{b - \frac{1}{2} - \mu}{\sigma} \right) - \Phi \left( \frac{a + \frac{1}{2} - \mu}{\sigma} \right).$$